Variance Component Models for Quantitative Traits

Biostatistics 666
Today

- Analysis of quantitative traits
- Kinship coefficients
  - measure of genetic similarity between two individuals
- Modeling covariance for pairs of individuals
  - estimating heritability
  - estimating locus-specific heritability
- Extending the model to larger pedigrees
Kinship Coefficients

- Summarize genetic similarity between pairs of individuals.
- In a variance components model, they predict the phenotypic similarity between individuals.
Kinship Coefficients – Definition

- Given two individuals
  - One with genes \((g_i, g_j)\)
  - The other with genes \((g_k, g_l)\)

- The kinship coefficient is:
  - \(\frac{1}{4}P(g_i=g_k) + \frac{1}{4}P(g_i=g_l) + \frac{1}{4}P(g_j=g_k) + \frac{1}{4}P(g_j=g_l)\)
  - where “\(\equiv\)” represents identity by descent (IBD)

- Probability that alleles sampled at random from each individual are IBD
Some kinship coefficients

- Siblings ($\varphi=1/4$)
- Half-Sibs ($\varphi=1/8$)
- MZ Twins ($\varphi=1/2$)
- Unrelated ($\varphi=0$)
- Parent-Offspring ($\varphi=1/4$)
What about other relatives?

- For any two related individuals i and j ...

- ... use a recursive algorithm allows calculation of kinship coefficient

- Algorithm requires an order for individuals in the pedigree where ancestors precede descendants
  - That is where for any $i > j$, $i$ is not ancestor of $j$
  - Such an order always exists (e.g. the birth order!)
Computing Kinship Coefficients

The recursive definition is then (for \( i \geq j \)):

\[
\phi_{ij} = \begin{cases} 
0 & \text{if } i \text{ and } j \text{ are founders} \\
\frac{1}{2} & \text{if } i = j, i \text{ is a founder} \\
\frac{1}{2} \left( \phi_{\text{mother}(i)j} + \phi_{\text{father}(i)j} \right) & \text{if } i \neq j \\
\frac{1}{2} \left( 1 + \phi_{\text{mother}(i)\text{father}(i)} \right) & \text{if } i = j
\end{cases}
\]
An example pedigree...

- Can you find ...
- Suitable ordering for recursive calculation?
- Calculate kinship coefficient between shaded individuals?
Inbreeding Coefficients

- The kinship coefficient is related to the inbreeding coefficient

- If $\phi_{ii} > 0.5$, individual $i$ is inbred

- The inbreeding coefficient is
  \[ f_i = \phi_{mother(i)father(i)} = 2(\phi_{ii} - 0.5) \]
So far ...

- Summarize genetic similarity between any two individuals ...

- Next, we will proceed to build a simple model for their phenotypes
Simplest Data Structure

- Pairs of related individuals
  - Siblings (or twins!)
  - Parent-Offspring

- Corresponding phenotype measurements
  - $y = (y_1, y_2)'$
If the trait is normally distributed …

Model mean and variance for $y_1$ and $y_2$
- Mean and variance could be assumed equal …
- … or they could depend on some covariates

But we are also interested in covariance between the two …
Variance-Covariance Matrix

\[
\Omega = \begin{bmatrix}
V(y_1) & Cov(y_1, y_2) \\
Cov(y_1, y_2) & V(y_2)
\end{bmatrix}
\]

Model must describe not only variance of each observation but also covariance for pairs of observations.
Bivariate density function

- Normal density function
  \[ L(y) = \frac{1}{\sqrt{2\pi}} \sigma^{-1} e^{-\frac{1}{2}(y-\mu)^2/\sigma^2} \]

- Bivariate normal density function
  \[ L(y) = \frac{1}{2\pi |\Omega|^{-1/2}} e^{-\frac{1}{2}(y-\mu)'\Omega^{-1}(y-\mu)} \]

- Extends univariate density function
Intuition on Normal Densities

\[ L(y) = (2\pi)^{-\frac{1}{2}} \sigma^{-1} e^{-\frac{1}{2} \left( \frac{(y-\mu)^2}{\sigma^2} \right)} \]

Scaling parameter, penalizes settings with large variances

Distance between observation and its expected value
Bivariate Normal Densities

\[ L(y) = \frac{1}{(2\pi)^{n/2} |\Omega|^{1/2}} e^{-\frac{1}{2} (y-\mu)'\Omega^{-1}(y-\mu)} \]

Scaling parameter, penalizes settings with large variances

Distance between observation and its expected value
Variability in Height, Independent Observations

![Histogram showing the distribution of height in centimeters. The x-axis represents height in centimeters ranging from 120 to 200, and the y-axis represents the number of individuals ranging from 0 to 400. The distribution is skewed with a peak around 160 centimeters.]
Variability in Height, Pairs of Observations

In a sample of twin or sibling pairs, we could use all the data to estimate means, variances and even covariances...

(Data from David Duffy)
Height in DZ and MZ twins

MZ twin Height \( r=0.96 \)

DZ twin Height \( r=0.32 \)

(How would you interpret these data from David Duffy?)
Incorporating Kinship Coefficients

- If genes influence trait …

- Covariance will differ for each class of relative pair

- Instead of estimating covariance for each relationship, …

- Impose genetic model that incorporates kinship and relates covariance between different classes of relative pair
A Simple Model for the Variance-Covariance Matrix

\[ \Omega = \begin{bmatrix} \sigma_g^2 + \sigma_e^2 & 2\varphi\sigma_g^2 \\ 2\varphi\sigma_g^2 & \sigma_g^2 + \sigma_e^2 \end{bmatrix} \]

Where,

\( \varphi \) is the kinship coefficient for the two individuals
(Reading ability scores from Eaves et al., 1997)
Interpretation...

- Fitting a maximum likelihood model...
  - Eaves et. al estimated
    - $\sigma_g^2 = .81$
    - $\sigma_e^2 = .19$
  - Found no evidence for sex differences
  - Saturated model did not improve fit
So far ...

- Model allows us to estimate the genetic contribution to the variation in any trait

- Incorporates different relative pairs ...

- But it doesn’t always fit…
  - Fortunately, the model can be easily refined
Another Example...

<table>
<thead>
<tr>
<th>Category</th>
<th>N</th>
<th>r</th>
</tr>
</thead>
<tbody>
<tr>
<td>MZ males</td>
<td>271</td>
<td>.56</td>
</tr>
<tr>
<td>MZ females</td>
<td>353</td>
<td>.52</td>
</tr>
<tr>
<td>DZ males</td>
<td>167</td>
<td>.33</td>
</tr>
<tr>
<td>DZ females</td>
<td>165</td>
<td>.45</td>
</tr>
<tr>
<td>DZ male-female</td>
<td>260</td>
<td>.41</td>
</tr>
</tbody>
</table>

(Psychomotor retardation scores from Eaves et al., 1997)
Refined Matrix

\[
\Omega = \begin{bmatrix}
\sigma_g^2 + \sigma_c^2 + \sigma_e^2 & 2\varphi\sigma_g^2 + \sigma_c^2 \\
2\varphi\sigma_g^2 + \sigma_c^2 & \sigma_g^2 + \sigma_c^2 + \sigma_e^2 \\
\end{bmatrix}
\]

Where,

\( \varphi \) is the kinship coefficient for the two individuals
Interpretation...

- Fitting a maximum likelihood model...
  - Eaves et. al estimated (for males)
    - $\sigma_g^2 = .29$
    - $\sigma_c^2 = .24$
    - $\sigma_e^2 = .46$
  - Additive genetic effects could not explain similarities. Any idea why?
Incorporating IBD Coefficients

- IBD coefficients measure genetic similarity at a specific locus
  - Related individuals might share 0, 1 or 2 alleles

- Covariance might differ according to sharing at a particular locus
  - If locus contains genes that influence the trait

- Again, impose a genetic model and estimate model parameters
Linkage

IBD 0

IBD 1

IBD 2
No Linkage

IBD 0

IBD 1

IBD 2
Relationship to IBD probabilities

For non-inbred pair of relatives, marker or locus-specific kinship coefficients can be derived from IBD probabilities:

\[
\varphi_{\text{marker}} = \frac{1}{4} P(\text{IBD}_{\text{marker}} = 1) + \frac{1}{2} P(\text{IBD}_{\text{marker}} = 2)
\]
Variance-Covariance Matrix

\[
\Omega = \begin{bmatrix}
\sigma_a^2 + \sigma_g^2 + \sigma_e^2 & 2\varphi_{\text{marker}}\sigma_a^2 + 2\varphi\sigma_g^2 \\
2\varphi_{\text{marker}}\sigma_a^2 + 2\varphi\sigma_g^2 & \sigma_a^2 + \sigma_g^2 + \sigma_e^2
\end{bmatrix}
\]

Where,

\( \varphi \) is the kinship coefficient for the two individuals

\( \varphi_{\text{marker}} \) depends on the number of alleles shared IBD
Likelihood function, Incorporating Uncertain IBD

\[ L = \prod_{i} \sum_{j=0,1,2} Z_{ij} (2\pi)^{-1} \mid \Omega_{IBD=j} \mid^{-\frac{1}{2}} e^{-\frac{1}{2}(y-\mu)'\Omega_{IBD=j}^{-1}(y-\mu)} \]

\[ \approx \prod_{i} (2\pi)^{-1} \mid \Omega^* \mid^{-\frac{1}{2}} e^{-\frac{1}{2}(y-\mu)'\Omega^{*-1}(y-\mu)} \]

\[ Z_{ij} = P(IBD_i = j \mid \text{marker data}) \quad \text{IBD sharing probabilities} \]

\[ \Omega^* = \sum_{j=0,1,2} Z_{ij} \Omega_{IBD=j} \]

"Expected" \( \Omega \)
How it works ...

- To find linkage to a particular trait…

- Collect sibling pair sample

- Calculate IBD for multiple points along genome

- Model covariance as a function of IBD sharing at each point
Example...

Estimated Major Gene Component $\sigma_a^2$
Likelihood Ratio Chisquared

\[ 2 \ln \frac{L(\sigma_a^2)}{L(\sigma_a^2=0)} \]
Example...

LOD Score

$$\log_{10} \frac{L(\sigma_a^2)}{L(\sigma_a^2=0)}$$
So far ...

- Models for similarity between relative pairs
- Kinship coefficient used to estimate overall genetic effect
- Locus-specific coefficients used to detect genetic linkage
Examples of Useful Extensions...

- For larger pedigrees, expand the covariance matrix.

- For populations with individuals of uncertain relatedness, estimate kinship using genotypes.

- To model genetic association, allow individual specific means to depend on genotype.
Larger Pedigrees...

\[ \Omega_{jk} = \begin{cases} 
\sigma_a^2 + \sigma_g^2 + \sigma_e^2 & \text{if } j = k \\
2\varphi_{\text{marker}}\sigma_a^2 + 2\varphi\sigma_g^2 & \text{if } j \neq k 
\end{cases} \]

Where,

\( \varphi \) is the kinship coefficient for the two individuals

\( \varphi_{\text{marker}} \) depends on the number of alleles shared IBD

\( j \) and \( k \) index different individuals in the family
Multivariate density function

- Normal density function

\[
L(y) = (2\pi)^{-\frac{1}{2}} \sigma^{-1} e^{-\frac{1}{2}(y-\mu)^2 / \sigma^2}
\]

- Multivariate normal density function

\[
L(y) = 2\pi^{-n/2} |\Omega|^{-1/2} e^{-\frac{1}{2}(y-\mu)'\Omega^{-1}(y-\mu)}
\]

- Extends univariate density function
Covariate and Genotype effects

- **Expected Phenotype for Individual** $i$ (e.g. expected weight)
- **Estimated effects for covariates** (e.g. expected weight increases 1kg/year with age)
- **Measured Covariates for Individual** $i$ (e.g. age, sex, genotype)

$$E(y_i) = \mu + \beta_1 x_{i1} + \beta_2 x_{i2} + \ldots + \beta_k x_{ik}$$

In addition to modeling variances and covariances, can model fixed effects.
Simple Association Model

- Each copy of allele changes trait by a fixed amount
  - Include covariate counting copies for allele of interest
- Evidence for association when \( a \neq 0 \)

\[ g_i = \text{number of copies of allele of interest in individual } i \]

\[ E(y_i) = \mu + \beta_g g_i \]

\( \beta_g \) is effect of each allele (the additive genetic value).
Relatedness in Populations

- Although we have focused on individuals of known relationship, ...

- Marker data can also be used to estimate relatedness.

- For example, Kang et al (2010) use:

\[ \hat{\phi}_{ij} = \frac{1}{M} \sum_{m=1}^{M} \frac{(g_{im} - 2p_m)(g_{jm} - 2p_m)}{4p_m(1 - p_m)} \]
Today

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  - Measure of genetic similarity between two individuals

- Modeling covariance for pairs of individuals
  - estimating heritability
  - estimating locus-specific heritability

- Extending the model to larger pedigrees
Useful References

- Amos (1994)
  Am J Hum Genet 54:535-543

- Hopper and Matthews (1982)
  Ann Hum Genet 46:373–383

- Lange and Boehnke (1983)