Biostatistics 602 - Statistical Inference Lecture 10 Maximum Likelihood Estimator

Hyun Min Kang

February 12th, 2013

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- 2 What is a method of moment estimator?

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- What is a method of moment estimator?
- What are advantages and disadvantages of method of moment estimator?
- 4 What is a maximum likelihood estimator (MLE)?
- 6 How can you find an MLE?

Recap - Method of Moment Estimator

- Point Estimation Estimate θ or $\tau(\theta)$.
- Method of Moment

$$m_1 = \frac{1}{n} \sum X_i = E\mathbf{X} = \mu_1$$

$$m_2 = \frac{1}{n} \sum X_i^2 = E\mathbf{X}^2 = \mu_2$$

$$\vdots$$

$$m_k = \frac{1}{n} \sum X_i^k = E\mathbf{X}^k = \mu_k$$

$$X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(\mu, \sigma^2)$$

$$\hat{\mu} = \overline{X}$$

$$\hat{\mu}^2 + \hat{\sigma}^2 = E\mathbf{X}^2 = \frac{1}{n} \sum_{i=1}^n X_i^2$$

$$\hat{\sigma}^2 = \sum (X_i - \overline{X})^2 / n$$

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- No guarantee that the estimator will fall into the range of valid parameter space.

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Recap - Likelihood Function

Definition

 $X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} f_X(x|\theta)$. The join distribution of $\mathbf{X} = (X_1, \dots, X_n)$ is

$$f_{\mathbf{X}}(\mathbf{x}|\theta) = \prod_{i=1}^{n} f_{X}(x_{i}|\theta)$$

Recap - Likelihood Function

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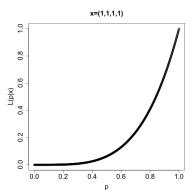
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Given that $\mathbf{X}=\mathbf{x}$ is observed, the function of θ defined by $L(\theta|\mathbf{x})=f(\mathbf{x}|\theta)$ is called the likelihood function.

Recap - Example Likelihood Function

- $X_1, X_2, X_3, X_4 \stackrel{\text{i.i.d.}}{\sim} \text{Bernoulli}(p), \ 0$
- $\mathbf{x} = (1, 1, 1, 1)^T$
- Intuitively, it is more likely that p is larger than smaller.
- $L(p|\mathbf{x}) = f(\mathbf{x}|p) = \prod_{i=1}^4 p^{x_i} (1-p)^{1-x_i} = p^4$.



If the function is differentiable with respect to θ ,

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(a)
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 or $\partial^2 L(\theta_1, \theta_2)^2 / \partial \theta_2^2 < 0$.

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 - (a) $\partial^2 L(\theta_1, \theta_2)^2 / \partial \theta_1^2 < 0$ or $\partial^2 L(\theta_1, \theta_2)^2 / \partial \theta_2^2 < 0$.
 - (b) Determinant of second-order derivative is positive

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- Or perform direct maximization, using inequalities, or properties of the function.

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Problem

 $X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\smile} \text{Uniform}(0, \theta)$, where $X_i \in [0, \theta]$ and $\theta > 0$.

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Solution

$$L(\theta|\mathbf{x}) = \prod_{i=1}^{n} \frac{1}{\theta} I(0 \le x_i \le \theta) = \frac{1}{\theta^n} \prod_{i=1}^{n} I(0 \le x_i \le \theta)$$

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We need to maximize $1/\theta^n$ subject to constraint that $0 \le x_{(n)} \le \theta$.

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Problem

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We need to maximize $1/\theta^n$ subject to constraint that $0 \le x_{(n)} \le \theta$. Because $1/\theta^n$ decreases in θ , the MLE is $\hat{\theta}(\mathbf{X}) = X_{(n)}$.

Problem

Suppose n pairs of data $(X_1, Y_1), \cdots, (X_n, Y_n)$ where X_i is generated from an unknown distribution, and Y_i are generated conditionally on X_i . $Y_i | X_i \sim \mathcal{N}(\alpha + \beta X_i, \sigma^2)$

Find the MLE of $(\alpha, \beta, \sigma^2)$.

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The joint distribution of $(X_1, Y_1), \dots, (X_n, Y_n)$ is

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$$\begin{aligned} f_{\mathbf{XY}}(\mathbf{x}, \mathbf{y}) &= f_{\mathbf{X}}(\mathbf{x}) \prod_{i=1}^{n} f_{\mathbf{Y}}(y_i | x_i) \\ &= f_{\mathbf{X}}(\mathbf{x}) \prod_{i=1}^{n} \frac{1}{2\pi\sigma^2} \exp \left[-\frac{(y_i - \alpha - \beta x_i)^2}{2\sigma^2} \right] \end{aligned}$$

The likelihood function is

$$L(\alpha, \beta, \sigma^2 | \mathbf{x}, \mathbf{y}) = f_{\mathbf{X}}(\mathbf{x})(2\pi\sigma^2)^{-n/2} \exp\left[-\frac{\sum_{i=1}^n (y_i - \alpha - \beta x_i)^2}{2\sigma^2}\right]$$

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$$l(\alpha, \beta, \sigma^2) = C - \frac{n}{2} \log(2\pi\sigma^2) - \frac{\sum_{i=1}^{n} (y_i - \alpha - \beta x_i)^2}{2\sigma^2}$$

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$$\frac{\partial l}{\partial \alpha} = \frac{2\sum_{i=1}^{n} (y_i - \alpha - \beta x_i)}{2\sigma^2} = \frac{n\overline{y} - n\alpha - n\beta \overline{x}}{\sigma^2} = 0$$

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$$\begin{array}{lcl} \frac{\partial l}{\partial \alpha} & = & \frac{2\sum_{i=1}^{n}(y_i-\alpha-\beta x_i)}{2\sigma^2} = \frac{n\overline{y}-n\alpha-n\beta\overline{x}}{\sigma^2} = 0 \\ \hat{\alpha} & = & \overline{y}-\hat{\beta}\overline{x} \end{array}$$

$$\frac{\partial l}{\partial \beta} = \frac{2\sum_{i=1}^{n} (y_i - \alpha - \beta x_i) x_i}{2\sigma^2} = \frac{\sum_{i=1}^{n} x_i y_i - n\alpha \overline{x} - \beta \sum_{i=1}^{n} x_i^2}{\sigma^2} = 0$$

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Putting Things Together

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$$\hat{\alpha} = \overline{y} - \hat{\beta}\overline{x}$$

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 $X_1, \cdots, X_n \overset{\text{i.i.d.}}{\smile} \mathcal{N}(\mu, 1)$ where $\mu \geq 0$. Find MLE of μ .

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$$L(\mu|\mathbf{x}) = \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{(x_i - \mu)^2}{2}\right] = (2\pi)^{-n/2} \exp\left[-\frac{\sum_{i=1}^{n} (x_i - \mu)^2}{2}\right]$$

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$$\hat{\mu} = \sum_{i=1}^{n} x_i / n = \overline{x}$$

Arawa dana?

Hyun Min Kang

We need to check whether $\hat{\mu}$ is within the parameter space $[0,\infty)$.

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• If $\overline{x} \geq 0$, $\hat{\mu} = \overline{x}$ falls into the parameter space.

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$$\hat{\mu}(\mathbf{X}) = \max(\overline{X}, 0)$$



Invariance Property of MLE

Question

If $\hat{\theta}$ is the MLE of θ , what is the MLE of $\tau(\theta)$?

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 \bullet What is the MLE of p?

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- \bullet What is the MLE of p?
- 2 What is the MLE of odds, defined by $\eta = p/(1-p)$?

$$L(p|\mathbf{x}) = \prod_{i=1}^{n} p^{x_i} (1-p)^{1-x_i} = p^{\sum x_i} (1-p)^{n-\sum x_i}$$

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$$\begin{split} L^*(\eta | \mathbf{x}) &= p^{\sum x_i} (1 - p)^{n - \sum x_i} \\ &= \frac{p}{1 - p}^{\sum x_i} (1 - p)^n = \frac{\eta^{\sum x_i}}{(1 + \eta)^n} \\ l^*(\eta | \mathbf{x}) &= \sum_{i=1}^n x_i \log \eta - n \log (1 + \eta) \\ \frac{\partial l^*}{\partial n} &= \frac{\sum_{i=1}^n x_i}{n} - \frac{n}{1 + n} = 0 \end{split}$$

MLE of $\eta = \frac{p}{1-p}$

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- Therefore $\hat{\eta} = \tau(\hat{p})$.

Fact

Denote the MLE of θ by $\hat{\theta}$. If $\tau(\theta)$ is an one-to-one function of θ , then MLE of $\tau(\theta)$ is $\tau(\hat{\theta})$.

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Proof

The likelihood function in terms of $\tau(\theta) = \eta$ is

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We know this function is maximized when $\tau^{-1}(\eta) = \hat{\theta}$, or equivalently, when $\eta = \tau(\hat{\theta})$. Therefore, MLE of $\eta = \tau(\theta)$ is $\tau(\hat{\theta})$.

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Summary

Today

Maximum Likelihood Estimator

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Maximum Likelihood Estimator

Next Lecture

- Mean Squared Error
- Unbiased Estimator
- Cramer-Rao inequality