

Biostatistics 602 - Statistical Inference

Lecture 14

Obtaining Best Unbiased Estimator

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Last Lecture

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- What is the Rao-Blackwell Theorem?
- Is the best unbiased estimator (UMVUE) for $\tau(\theta)$ is unique?
- What is the relationship between the UMVUE and the unbiased estimators of zero?

Rao-Blackwell Theorem

Theorem 7.3.17

Let $W(\mathbf{X})$ be any unbiased estimator of $\tau(\theta)$, and T be a sufficient statistic for θ . Define $\phi(T) = E[W|T]$. Then the followings hold.

- 1 $E[\phi(T)|\theta] = \tau(\theta)$
- 2 $\text{Var}[\phi(T)|\theta] \leq \text{Var}(W|\theta)$ for all θ .

That is, $\phi(T)$ is a uniformly better unbiased estimator of $\tau(\theta)$.

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Theorem 7.3.20 - UMVUE and unbiased estimators of zero

If $E[W(\mathbf{X})] = \tau(\theta)$. W is the best unbiased estimator of $\tau(\theta)$ if and only if W is uncorrelated with all unbiased estimator of 0.

The power of complete sufficient statistics

Theorem 7.3.23

Let T be a complete sufficient statistic for parameter θ . Let $\phi(T)$ be any estimator based on T . Then $\phi(T)$ is the unique best unbiased estimator of its expected value.

Remarks from previous Theorems - #1

From Rao-Blackwell Theorem, we can always improve an unbiased estimator by conditioning it on a sufficient statistics.

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- $W(\mathbf{X})$: unbiased for $\tau(\theta)$.
- $T^*(\mathbf{X})$: sufficient statistic for θ .

$\phi(T) = E[W(\mathbf{X}) | T(\mathbf{X})]$ is a better unbiased estimator of $\tau(\theta)$.

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In fact, we only need to consider functions of minimal sufficient statistics to find the best unbiased estimator.

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In fact, we only need to consider functions of minimal sufficient statistics to find the best unbiased estimator.

Let $T(\mathbf{X})$ be a minimal sufficient, and $T^*(\mathbf{X})$ be a sufficient statistic. Then by definition, there exists a function h that satisfies $T = h(T^*)$.

$$\begin{aligned} E[\phi(T) | T^*] &= E[\phi\{h(T^*)\} | T^*] \\ &= \phi\{h(T^*)\} = \phi(T) \end{aligned}$$

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$$\begin{aligned} E[\phi(T) | T^*] &= E[\phi\{h(T^*)\} | T^*] \\ &= \phi\{h(T^*)\} = \phi(T) \end{aligned}$$

Therefore $\phi(T)$ remains the same after conditioning on any sufficient statistic T^* .

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- Suppose that T is a complete statistic, then $U(T)$ can only be zero almost surely.

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- By definition, T is complete if $E[U(T)] = 0$ for all θ implies $U(T) = 0$ almost surely.
- Suppose that T is a complete statistic, then $U(T)$ can only be zero almost surely.
- Therefore, $\text{Cov}(\phi(T), U(T)) = \text{Cov}(\phi(T), 0) = 0$, and $\phi(T)$ is the best unbiased estimator of its expected value (Theorem 7.3.23).

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 - Guess a function $\phi(T)$ such that $E[\phi(T)] = \tau(\theta)$.
 - Guess an unbiased estimator $h(\mathbf{X})$ of $\tau(\theta)$. Construct $\phi(T) = E[h(\mathbf{X}) | T]$, then $E[\phi(T)] = E[h(\mathbf{X})] = \tau(\theta)$.

Example - Normal Distribution

Problem

$X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(\mu, \sigma^2)$. Find the best unbiased estimator for (1) μ , (2) σ^2 , (3) μ^2 .

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- Therefore, \bar{X} is the best unbiased estimator for μ .

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- Therefore $s_{\mathbf{X}}^2$ is the best unbiased estimator of σ^2 .

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- $\bar{X}^2 - s_{\mathbf{X}}^2/n$ is unbiased estimator for μ^2
- And it is a function of $(\bar{X}, s_{\mathbf{X}}^2)$.
- Hence, $\bar{X}^2 - s_{\mathbf{X}}^2/n$ is the best unbiased estimator for μ^2 .

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Example - Uniform Distribution

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$X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} \text{Uniform}(0, \theta)$. Find the best unbiased estimator for (1) θ , (2) $g(\theta)$ differentiable on $(0, \theta)$ (3) θ^2 , (4) $1/\theta$.

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- $E[\phi(T)] = E\left[\frac{n+1}{n}X_{(n)}\right] = \theta$.

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- $E[\phi(T)] = E\left[\frac{n+1}{n}X_{(n)}\right] = \theta$.

$\frac{n+1}{n}X_{(n)}$ is the best unbiased estimator of θ .

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We need to find a function of $\phi(T) = X_{(n)}$ such that $E[\phi(T)] = g(\theta)$.

$$g(\theta) = E[\phi(T)] = \int_0^\theta \phi(t) n \theta^{-n} t^{n-1} dt$$

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Taking derivative with respect to θ , and applying Leibnitz's rule.

$$g'(\theta) = \frac{d}{d\theta} \int_0^\theta \phi(t) n \theta^{-n} t^{n-1} dt$$

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Example - Uniform Distribution - for $g(\theta)$

We need to find a function of $\phi(T) = X_{(n)}$ such that $E[\phi(T)] = g(\theta)$.

$$g(\theta) = E[\phi(T)] = \int_0^\theta \phi(t) n \theta^{-n} t^{n-1} dt$$

Taking derivative with respect to θ , and applying Leibnitz's rule.

$$\begin{aligned} g'(\theta) &= \frac{d}{d\theta} \int_0^\theta \phi(t) n \theta^{-n} t^{n-1} dt \\ &= \phi(\theta) n \theta^{-n} \theta^{n-1} + \int_0^\theta \phi(t) t^{n-1} n \frac{d}{d\theta} \theta^{-n} dt \\ &= \phi(\theta) n \theta^{-1} + \int_0^\theta \phi(t) t^{n-1} n (-n) \theta^{-n-1} dt \\ &= \phi(\theta) n \theta^{-1} - n \theta^{-1} \int_0^\theta \phi(t) n t^{n-1} \theta^{-n} dt \end{aligned}$$

Example - Uniform Distribution - for $g(\theta)$ (cont'd)

$$g'(\theta) = \phi(\theta)n\theta^{-1} - n\theta^{-1} \int_0^\theta \phi(t)nt^{n-1}\theta^{-n}dt$$

Example - Uniform Distribution - for $g(\theta)$ (cont'd)

$$\begin{aligned} g'(\theta) &= \phi(\theta)n\theta^{-1} - n\theta^{-1} \int_0^\theta \phi(t)nt^{n-1}\theta^{-n}dt \\ &= \phi(\theta)n\theta^{-1} - n\theta^{-1}E[\phi(T)] \end{aligned}$$

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$$\begin{aligned}g'(\theta) &= \phi(\theta)n\theta^{-1} - n\theta^{-1} \int_0^\theta \phi(t)nt^{n-1}\theta^{-n}dt \\&= \phi(\theta)n\theta^{-1} - n\theta^{-1}E[\phi(T)] \\&= \phi(\theta)n\theta^{-1} - n\theta^{-1}g(\theta) \\ \phi(\theta) &= \frac{g'(\theta) + n\theta^{-1}g(\theta)}{n\theta^{-1}}\end{aligned}$$

Therefore, the best unbiased estimator of $g(\theta)$ is

$$\phi(T) = \frac{g'(T) + nT^{-1}g(T)}{nT^{-1}}$$

Example - Uniform Distribution - for $g(\theta)$ (cont'd)

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Therefore, the best unbiased estimator of $g(\theta)$ is

$$\begin{aligned}\phi(T) &= \frac{g'(T) + nT^{-1}g(T)}{nT^{-1}} \\ \phi(X_{(n)}) &= \frac{g'(X_{(n)}) + nX_{(n)}^{-1}g(X_{(n)})}{nX_{(n)}^{-1}} \\ &= \frac{1}{n}X_{(n)}g'(X_{(n)}) + g(X_{(n)})\end{aligned}$$

Example - Uniform Distribution - for θ^2

$$g(\theta) = \theta^2, \text{ and } g'(\theta) = 2\theta.$$

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$g(\theta) = \theta^2$, and $g'(\theta) = 2\theta$.

$$\begin{aligned}\phi(X_{(n)}) &= \frac{1}{n}X_{(n)} \cdot 2X_{(n)} + X_{(n)}^2 \\ &= \frac{n+2}{n}X_{(n)}^2\end{aligned}$$

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Example - Binomial best unbiased estimator

Problem

$X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} \text{Binomial}(k, \theta)$. Estimate the probability of exactly one success.

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$$\tau(\theta) = \Pr(X = 1|\theta) = k\theta(1 - \theta)^{k-1}$$

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- We know that $T(\mathbf{X}) = \sum_{i=1}^n X_i \sim \text{Binomial}(kn, \theta)$ and it is a complete sufficient statistic.
- So we need to find a $\phi(T)$ that satisfies $E[\phi(T)] = \tau(\theta)$.
- There is no immediately evident unbiased estimator of $\tau(\theta)$ as a function of T .

Solution - Binomial best unbiased estimator

- Start with a simple-minded estimator

$$W(X_1) = \begin{cases} 1 & X_1 = 1 \\ 0 & \text{otherwise} \end{cases}$$

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and hence is an unbiased estimator of $\tau(\theta) = k\theta(1-\theta)^{k-1}$.

- The best unbiased estimator of $\tau(\theta)$ is

$$\phi(T) = E[W|T] = E[W(X_1)|T(\mathbf{X})]$$

Solution - Binomial best unbiased estimator (cont'd)

$$\phi(t) = E \left[W(X_1) \mid \sum_{i=1}^n X_i = t \right] = \Pr \left[X_1 = 1 \mid \sum_{i=1}^n X_i = t \right]$$

Solution - Binomial best unbiased estimator (cont'd)

$$\begin{aligned}\phi(t) &= E \left[W(X_1) \mid \sum_{i=1}^n X_i = t \right] = \Pr \left[X_1 = 1 \mid \sum_{i=1}^n X_i = t \right] \\ &= \frac{\Pr(X_1 = 1, \sum_{i=1}^n X_i = t)}{\Pr(\sum_{i=1}^n X_i = t)}\end{aligned}$$

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Solution - Binomial best unbiased estimator (cont'd)

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Solution - Binomial best unbiased estimator (cont'd)

$$\begin{aligned}\phi(t) &= E \left[W(X_1) \mid \sum_{i=1}^n X_i = t \right] = \Pr \left[X_1 = 1 \mid \sum_{i=1}^n X_i = t \right] \\&= \frac{\Pr(X_1 = 1, \sum_{i=1}^n X_i = t)}{\Pr(\sum_{i=1}^n X_i)} \\&= \frac{\Pr(X_1 = 1, \sum_{i=2}^n X_i = t-1)}{\Pr(\sum_{i=1}^n X_i)} \\&= \frac{\Pr(X_1 = 1) \Pr(\sum_{i=2}^n X_i = t-1)}{\Pr(\sum_{i=1}^n X_i)} \\&= \frac{[k\theta(1-\theta)^{k-1}] \left[\binom{k(n-1)}{t-1} \theta^{t-1} (1-\theta)^{k(n-1)-t+1} \right]}{\binom{kn}{n} \theta^t (1-\theta)^{kn-t}}\end{aligned}$$

Solution - Binomial best unbiased estimator (cont'd)

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Solution - Binomial best unbiased estimator (cont'd)

Therefore, the unbiased estimator of $k\theta(1 - \theta)^{k-1}$ is

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Therefore, the unbiased estimator of $k\theta(1 - \theta)^{k-1}$ is

$$\phi\left(\sum_{i=1}^n X_i\right) = k \frac{\binom{k(n-1)}{\sum X_i - 1}}{\binom{kn}{\sum X_i}}$$

Summary

Today

- Rao-Blackwell Theorem
- Methods for obtaining UMVUE

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- Rao-Blackwell Theorem
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Next Lecture

- Bayesian Estimators