

*Multipoint Analysis for  
Sibling Pairs*

**Biostatistics 666**

## Previously ...

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- Linkage analysis with pairs of individuals
  - “Non-parametric” IBS Methods
  - “Maximum Likelihood” IBD Based Method
  - Possible triangle constraint
- Examined one marker at a time

## IBS Based Linkage Test

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$$\chi^2_{2df} = \sum_i \frac{[N_{IBS=i} - E(N_{IBS=i})]^2}{E(N_{IBS=i})}$$

- Expect counts calculated using:
  - Allele frequencies for marker
  - Relationship for affected individuals

## Likelihood for a Single ASP

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$$L_i = \sum_{j=0}^2 P(IBD = j | ASP) P(Genotypes | IBD = j) = \sum_{j=0}^2 z_j w_{ij}$$

Risch (1990) defines

$$w_{ij} = P(Genotypes_i | IBD = j)$$

We only need proportionate  $w_{ij}$

# MLS Linkage Test

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$$L(z_0, z_1, z_2) = \prod_i \sum_j z_j w_{ij}$$

$$LOD = \log_{10} \prod_i \frac{z_0 w_{i0} + z_1 w_{i1} + z_2 w_{i2}}{\frac{1}{4} w_{i0} + \frac{1}{2} w_{i1} + \frac{1}{4} w_{i2}}$$

**The MLS statistic is the LOD evaluated at the MLEs of  $z_0, z_1, z_2$**

## Important Limitation

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- A major limitation of all these models is that they consider only one marker at a time
- This may not allow us to extract all available information about IBD...

## Today ...

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- Refresher on IBD probabilities
- Intuition behind multipoint calculations
- Framework for multipoint calculations
- Using a Markov Chain to speed analyses

## IBD Probabilities

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- Number of alleles identical by descent
- For sibling pairs, must be:
  - 0
  - 1
  - 2
- Not always determined by marker data



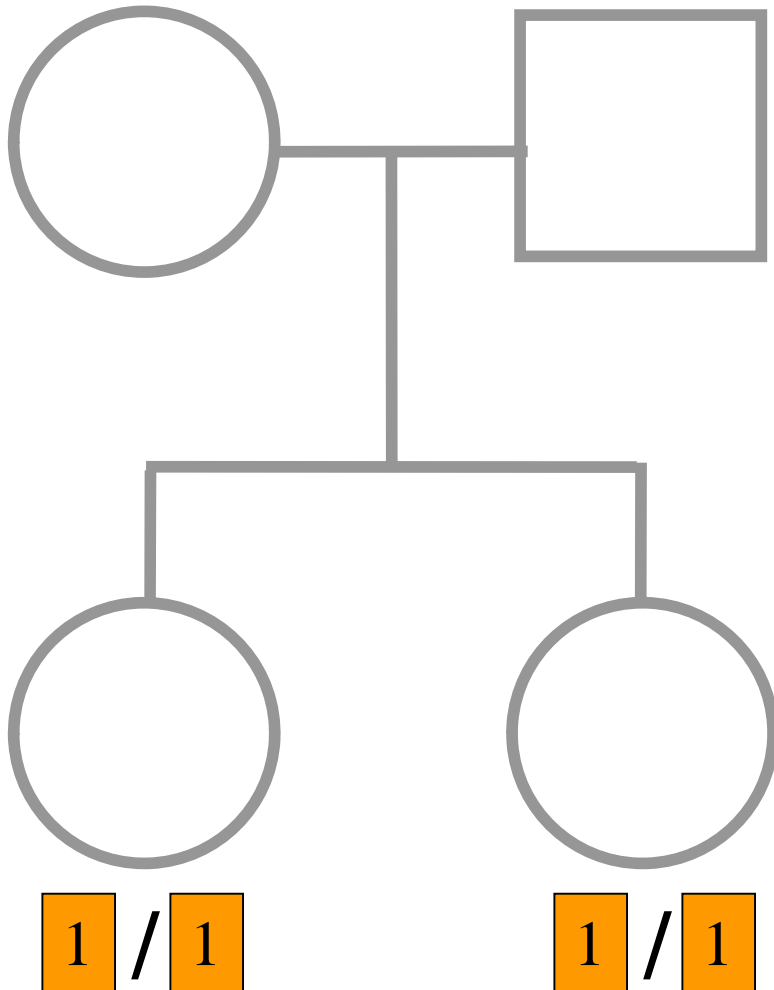
Refresher ...

## Bayes Theorem for IBD Probabilities

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$$\begin{aligned} P(IBD = i | X) &= \frac{P(IBD = i, X)}{P(X)} \\ &= \frac{P(IBD = i)P(X | IBD = i)}{P(X)} \\ &= \frac{P(IBD = i)P(X | IBD = i)}{\sum_j P(IBD = j)P(X | IBD = j)} \end{aligned}$$

## Worked Example



$$p_1 = 0.5$$

$$P(X | IBD=0) = p_1^4 = \frac{1}{16}$$

$$P(X | IBD=1) = p_1^3 = \frac{1}{8}$$

$$P(X | IBD=2) = p_1^2 = \frac{1}{4}$$

$$P(X) = \frac{1}{4}p_1^4 + \frac{1}{2}p_1^3 + \frac{1}{4}p_1^2 = \frac{9}{64}$$

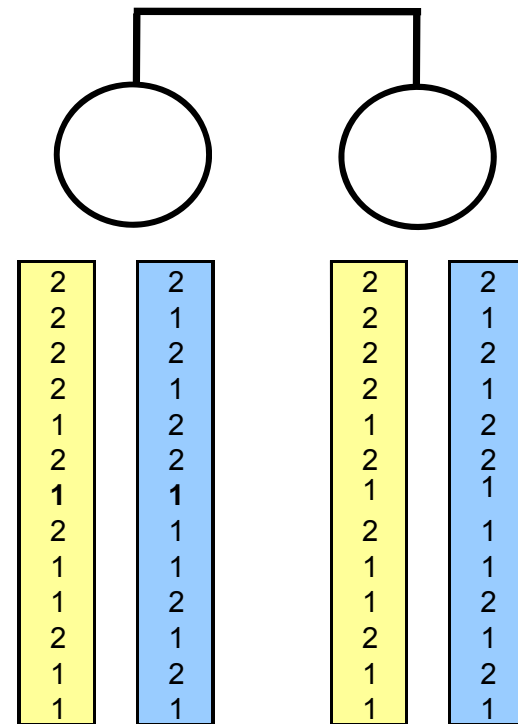
$$P(IBD=0 | X) = \frac{\frac{1}{4}p_1^4}{P(X)} = \frac{1}{9}$$

$$P(IBD=1 | X) = \frac{\frac{1}{2}p_1^3}{P(X)} = \frac{4}{9}$$

$$P(IBD=2 | X) = \frac{\frac{1}{4}p_1^2}{P(X)} = \frac{4}{9}$$

# Intuition For Multipoint Analysis

- IBD changes infrequently along the chromosome
- Neighboring markers can help resolve ambiguities about IBD sharing
- In the Risch approach, they might ensure that only one  $w$  is *effectively* non-zero



## The Ingredients ...

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- Probability of observed genotypes at each marker conditional on IBD state
- Probability of changes in IBD state along chromosome
- Hidden Markov Model

# Ingredients

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$X_1$

$X_2$

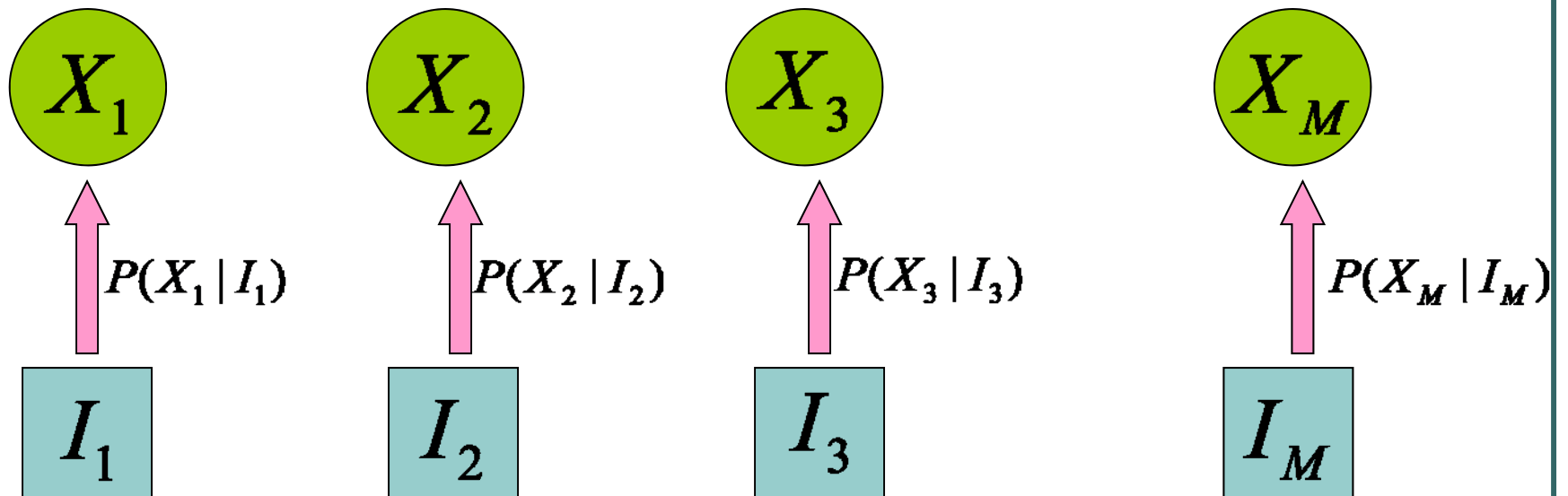
$X_3$

$X_M$

One ingredient will be the observed genotypes at each marker ...

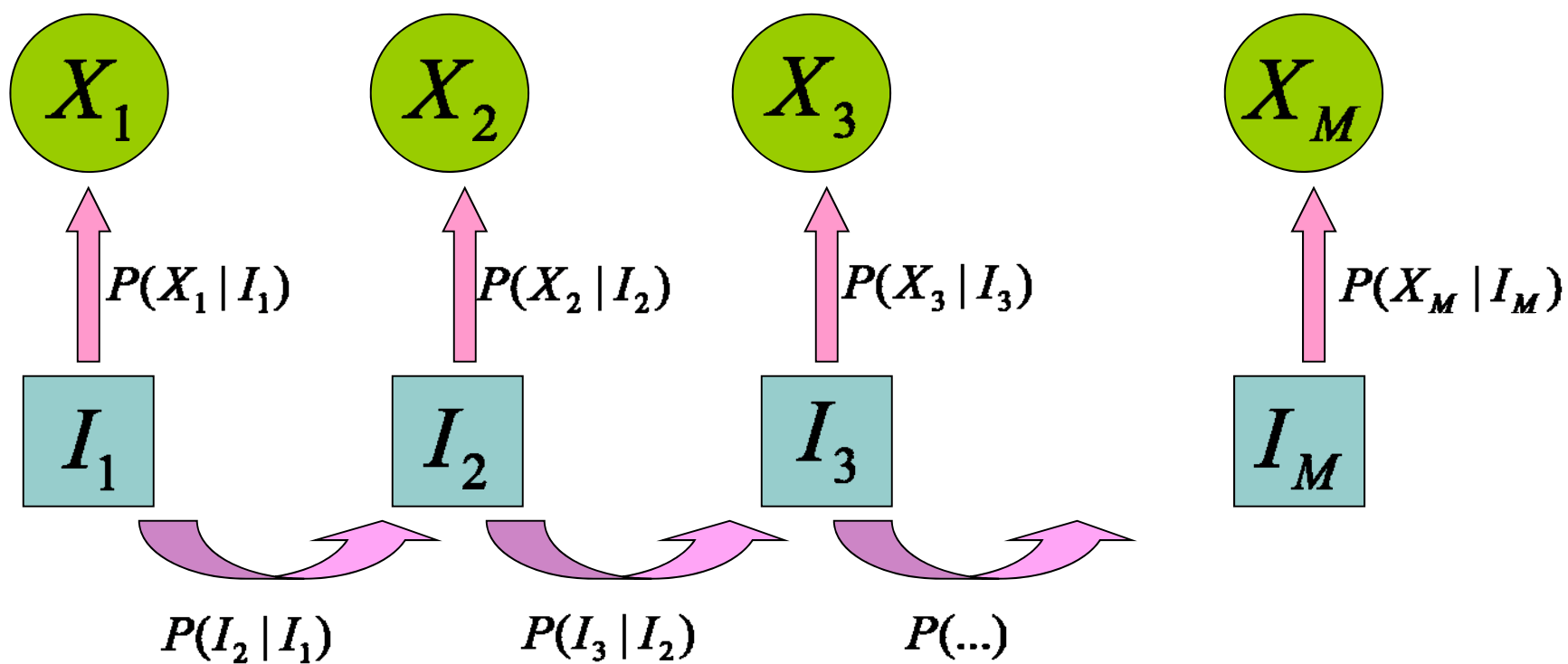
# Ingredients

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Another ingredient will be the possible IBD states at each marker ...

# Ingredients



The final ingredient connects IBD states along the chromosome ...

## The Likelihood of Marker Data

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$$L = \sum_{I_1} \sum_{I_2} \dots \sum_{I_M} P(I_1) \prod_{i=2}^M P(I_i | I_{i-1}) \prod_{i=1}^M P(X_i | I_i)$$

- General formulation, allows for any number of markers.
- Combined with Bayes' Theorem can estimate probability of each IBD state at any marker.
- This is not a linkage test yet!



$$P(X_m | I_m)$$

Sib	CoSib	IBD		
		0	1	2
(a,b)	(c,d)	$4p_a p_b p_c p_d$	0	0
(a,a)	(b,c)	$2p_a^2 p_b p_c$	0	0
(a,a)	(b,b)	$p_a^2 p_b^2$	0	0
(a,b)	(a,c)	$4p_a^2 p_b p_c$	$p_a p_b p_c$	0
(a,a)	(a,b)	$2p_a^3 p_b$	$p_a^2 p_b$	0
(a,b)	(a,b)	$4p_a^2 p_b^2$	$(p_a p_b^2 + p_a^2 p_b)$	$2p_a p_b$
(a,a)	(a,a)	$p_a^4$	$p_a^3$	$p_a^2$
<b>Prior Probability</b>		$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$

Note: Assuming unordered genotypes

Question:

What to do about missing data?

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- What happens when some genotype data is unavailable?

$$P(I_{m+1} | I_m)$$

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- Depends on recombination fraction  $\theta$ 
  - This is a measure of distance between two loci
  - Probability grand-parental origin of alleles changes between loci
- Naturally, leads to probability of change in IBD:

$$\psi = 2\theta(1 - \theta)$$

$$P(I_{m+1} | I_m)$$


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		IBD State at m + 1		
		0	1	2
IBD state at marker m	0	$(1-\psi)^2$	$2\psi(1-\psi)$	$\psi^2$
	1	$\psi(1-\psi)$	$(1-\psi)^2 + \psi^2$	$\psi(1-\psi)$
	2	$\psi^2$	$2\psi(1-\psi)$	$(1-\psi)^2$

$$\psi = 2\theta(1-\theta)$$

## Example

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- Consider two loci separated by  $\theta = 0.1$
- Each loci has two alleles, each with frequency .50
- If two siblings are homozygous for the first allele at both loci, what is the probability that IBD = 2 at the first locus?

## The Likelihood of Marker Data

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$$L = \sum_{I_1} \sum_{I_2} \dots \sum_{I_M} P(I_1) \prod_{i=2}^M P(I_i | I_{i-1}) \prod_{i=1}^M P(X_i | I_i)$$

- General, but slow unless there are only a few markers.
- How do we speed things up?

# A Markov Model

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- Re-organize the computation slightly, to avoid evaluating nested sum directly
- Three components:
  - Probability considering a single location
  - Probability including left flanking markers
  - Probability including right flanking markers
- Scale of computation increases linearly with number of markers

## The Likelihood of Marker Data

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$$\begin{aligned} L &= \sum_{I_j} P(I_j) P(X_j | I_j) P(X_1 \dots X_{j-1} | I_j) P(X_{j+1} \dots X_M | I_j) \\ &= \sum_{I_j} P(I_j) P(X_j | I_j) L_j(I_j) R_j(I_j) \end{aligned}$$

- A different arrangement of the same likelihood
- The nested summations are now hidden inside the  $L_j$  and  $R_j$  functions...



## Left-Chain Probabilities

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$$\begin{aligned} L_m(I_m) &= P(X_1, \dots, X_{m-1} | I_m) \\ &= \sum_{I_{m-1}} L_{m-1}(I_{m-1}) P(X_{m-1} | I_{m-1}) P(I_{m-1} | I_m) \end{aligned}$$

$$L_1(I_1) = 1$$

- Proceed one marker at a time.
- Computation cost increases linearly with number of markers.

## Right-Chain Probabilities

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$$\begin{aligned} R_m(I_m) &= P(X_{m+1}, \dots, X_M | I_m) \\ &= \sum_{I_{m+1}} R_{m+1}(I_{m+1}) P(X_{m+1} | I_{m+1}) P(I_{m+1} | I_m) \end{aligned}$$

$$R_M(I_M) = 1$$

- Proceed one marker at a time.
- Computation cost increases linearly with number of markers.

# Pictorial Representation

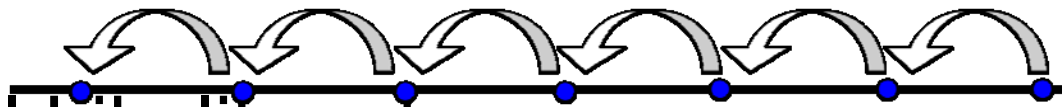
- Single Marker



- Left Conditional



- Right Conditional



- Full Likelihood



## Extending the MLS Method ...

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$$\begin{aligned}w_j &= P(X_j | I_j)P(X_1 \dots X_{j-1} | I_j)P(X_{j+1} \dots X_M | I_j) \\ &= P(X_j | I_j)L_j(I_j)R_j(I_j)\end{aligned}$$

- We just change the definition for the “weights” given to each configuration!

## Some Extensions We'll Discuss

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- Modeling error
  - What components might have to change?
- Modeling other types of relatives
  - What components might have to change?
- Modeling larger pedigrees
  - What components might have to change?

# Today

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- Efficient computational framework for multipoint analysis of sibling pairs