

*Variance Component Models
for Quantitative Traits*

Biostatistics 666

Today

- Analysis of quantitative traits
- Kinship coefficients
 - measure of genetic similarity between two individuals
- Modeling covariance for pairs of individuals
 - estimating heritability
 - estimating locus-specific heritability
- Extending the model to larger pedigrees

Kinship Coefficients

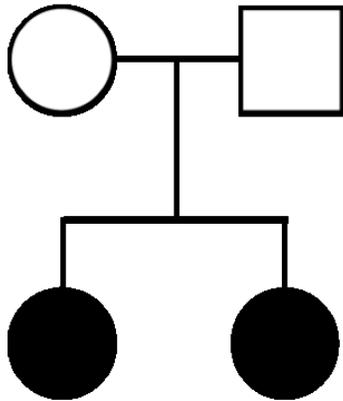
- Summarize genetic similarity between pairs of individuals.
- In a variance components model, they predict the phenotypic similarity between individuals.

Kinship Coefficients – Definition

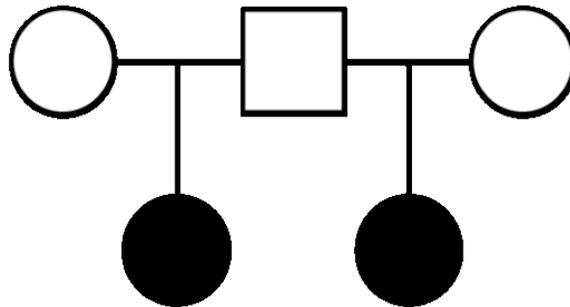
- Given two individuals
 - One with genes (g_i, g_j)
 - The other with genes (g_k, g_l)
- The kinship coefficient is:
 - $\frac{1}{4}P(g_i \equiv g_k) + \frac{1}{4}P(g_i \equiv g_l) + \frac{1}{4}P(g_j \equiv g_k) + \frac{1}{4}P(g_j \equiv g_l)$
 - where “ \equiv ” represents identity by descent (IBD)
- Probability that alleles sampled at random from each individual are (IBD)

Some kinship coefficients

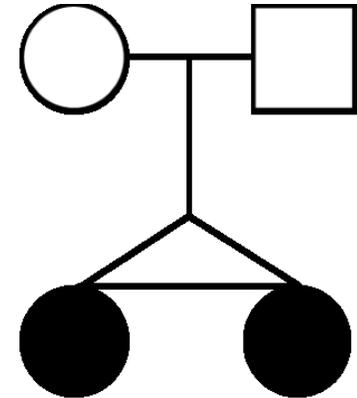
Siblings ($\phi=1/4$)



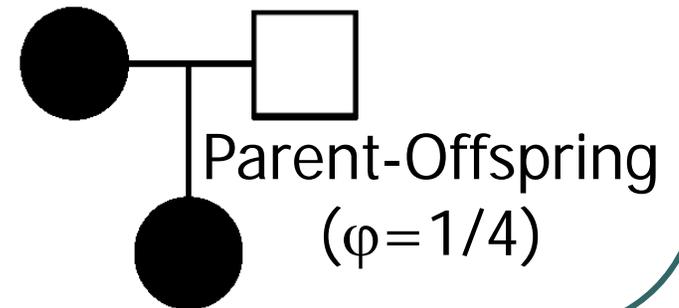
Half-Sibs ($\phi=1/8$)



MZ Twins ($\phi=1/2$)



Unrelated ($\phi=0$)



Parent-Offspring
($\phi=1/4$)

What about other relatives?

- For any two related individuals i and j ...
- ... use a recursive algorithm allows calculation of kinship coefficient
- Algorithm requires an order for individuals in the pedigree where ancestors precede descendants
 - That is where for any $i > j$, i is not ancestor of j
 - Such an order always exists!

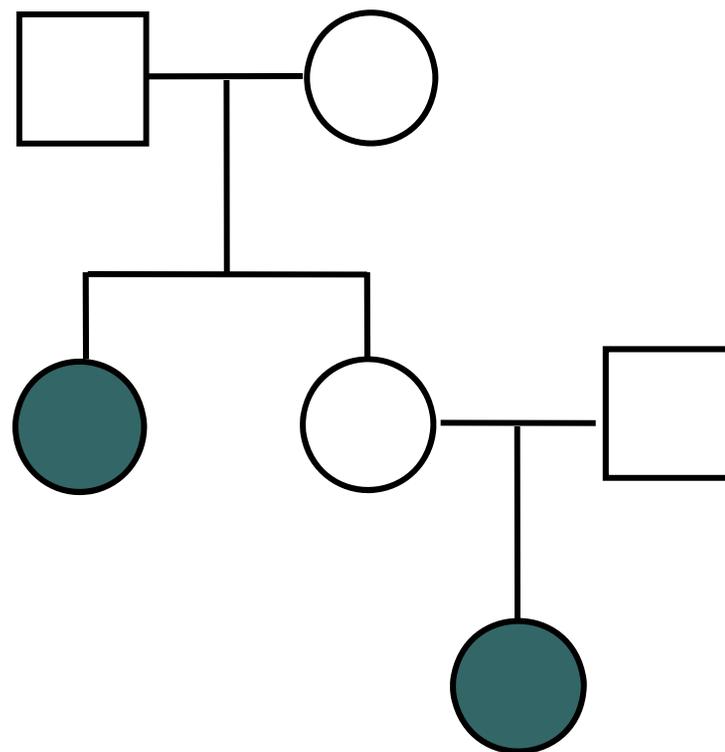
Computing Kinship Coefficients

- The recursive definition is then (for $i \geq j$):

$$\varphi_{ij} = \begin{cases} 0 & i \text{ and } j \text{ are founders} \\ \frac{1}{2} & i = j, i \text{ is a founder} \\ \frac{1}{2}(\varphi_{\text{mother}(i)j} + \varphi_{\text{father}(i)j}) & i \neq j \\ \frac{1}{2}(1 + \varphi_{\text{mother}(i)\text{father}(i)}) & i = j \end{cases}$$

An example pedigree...

- Can you find ...
- Suitable ordering for recursive calculation?
- Calculate kinship coefficient between shaded individuals?



So far ...

- Summarize genetic similarity between any two individuals ...
- Next, we will proceed to build a simple model for their phenotypes

Simplest Data Structure

- Pairs of related individuals
 - Siblings (or twins!)
 - Parent-Offspring
- Corresponding phenotype measurements
 - $\mathbf{y} = (\mathbf{y}_1, \mathbf{y}_2)'$

Elements for a simple model ...

- If the trait is normally distributed ...
- Model mean and variance for \mathbf{y}_1 and \mathbf{y}_2
 - Mean and variance could be assumed equal ...
 - ... or they could depend on some covariates
- But we are also interested in covariance between the two ...

Variance-Covariance Matrix

$$\Omega = \begin{bmatrix} V(y_1) & Cov(y_1, y_2) \\ Cov(y_1, y_2) & V(y_2) \end{bmatrix}$$

Model must describe not only variance of each observation but also covariance for pairs of observations

Bivariate density function

- Normal density function

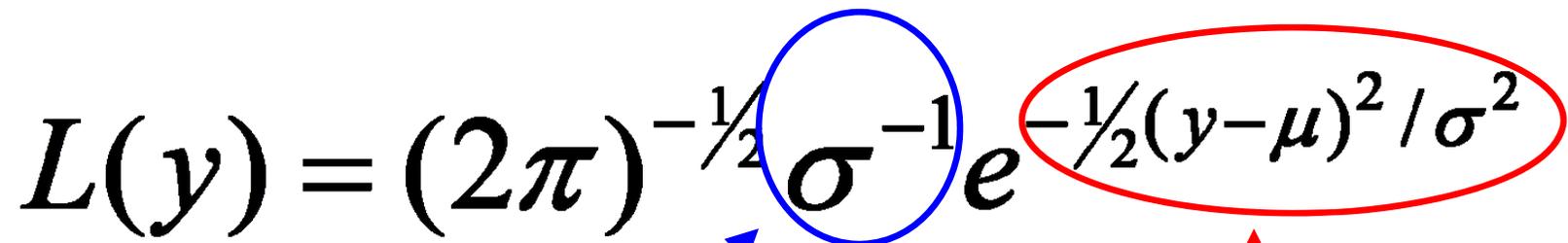
$$L(y) = \frac{1}{\sqrt{2\pi}} \sigma^{-1} e^{-\frac{1}{2}(y-\mu)^2 / \sigma^2}$$

- Bivariate normal density function

$$L(\mathbf{y}) = \frac{1}{2\pi} |\mathbf{\Omega}|^{-1/2} e^{-\frac{1}{2}(\mathbf{y}-\boldsymbol{\mu})' \mathbf{\Omega}^{-1} (\mathbf{y}-\boldsymbol{\mu})}$$

- Extends univariate density function

Intuition on Normal Densities

$$L(y) = (2\pi)^{-1/2} \sigma^{-1} e^{-1/2(y-\mu)^2 / \sigma^2}$$


Scaling parameter, penalizes settings with large variances

Distance between observation and its expected value

Bivariate Normal Densities

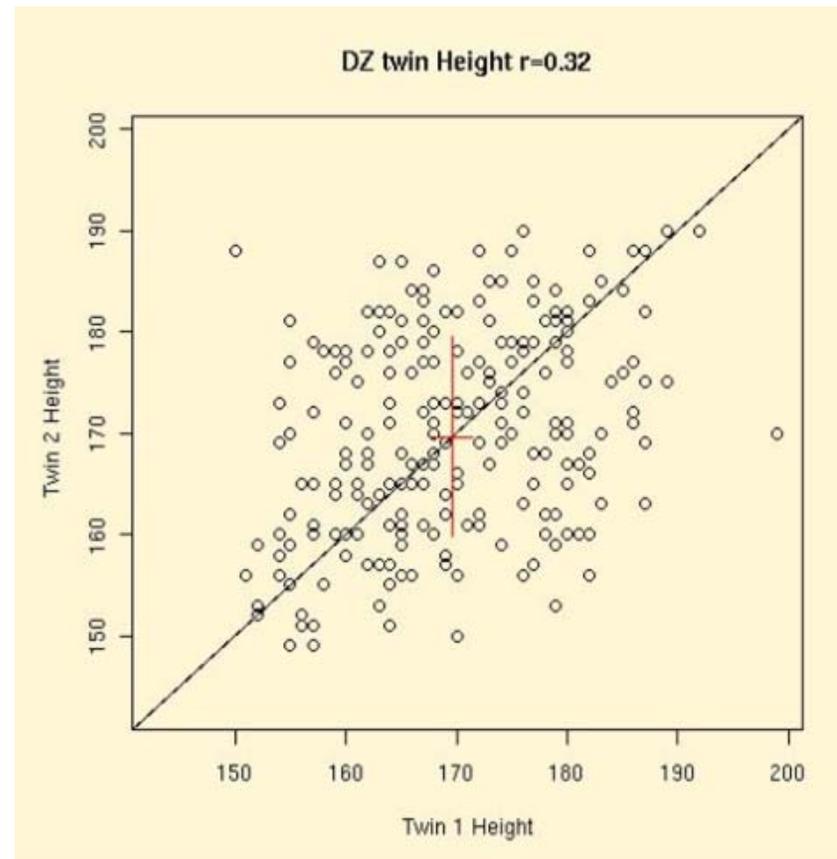
$$L(\mathbf{y}) = (2\pi)^{-1} |\mathbf{\Omega}|^{-1/2} e^{-1/2(\mathbf{y}-\boldsymbol{\mu})'\mathbf{\Omega}^{-1}(\mathbf{y}-\boldsymbol{\mu})}$$

Scaling parameter, penalizes settings with large variances

Distance between observation and its expected value

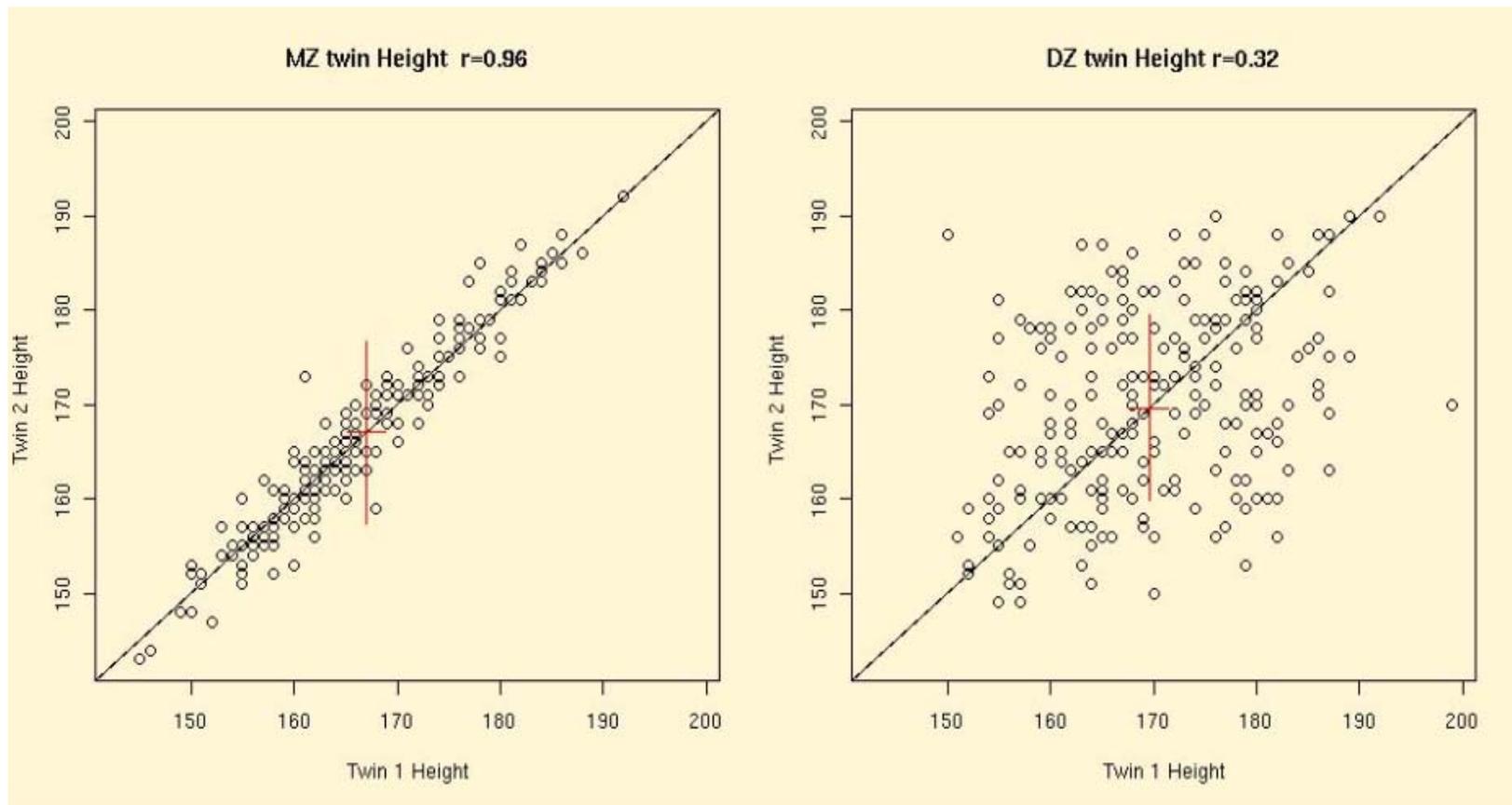
Possible Application...

In a sample of twin or sibling pairs, we could use all the data to estimate means, variances and even covariances...



(Data from David Duffy)⁶

Height in DZ and MZ twins



(How would you interpret these data from David Duffy?)

Incorporating Kinship Coefficients

- If genes influence trait ...
- Covariance will differ for each class of relative pair
- Instead of estimating covariance for each relationship, ...
- Impose genetic model that incorporates kinship and relates covariance between different classes of relative pair

A Simple Model for the Variance-Covariance Matrix

$$\Omega = \begin{bmatrix} \sigma_g^2 + \sigma_e^2 & 2\varphi\sigma_g^2 \\ 2\varphi\sigma_g^2 & \sigma_g^2 + \sigma_e^2 \end{bmatrix}$$

Where,

φ is the kinship coefficient for the two individuals

Example...

| | N | r |
|----------------|----------|----------|
| MZ males | 292 | .80 |
| MZ females | 380 | .80 |
| DZ males | 179 | .47 |
| DZ females | 184 | .55 |
| DZ male-female | 284 | .41 |

(Reading ability scores from Eaves et al., 1997)

Interpretation...

- Fitting a maximum likelihood model...
 - Eaves et. al estimated
 - $\sigma_g^2 = .81$
 - $\sigma_e^2 = .19$
 - Found no evidence for sex differences
 - Saturated model did not improve fit

So far ...

- Model allows us to estimate the genetic contribution to the variation in any trait
- Incorporates different relative pairs ...
- But it doesn't always fit...
 - Fortunately, the model can be easily refined

Another Example...

| | N | r |
|----------------|----------|----------|
| MZ males | 271 | .56 |
| MZ females | 353 | .52 |
| DZ males | 167 | .33 |
| DZ females | 165 | .45 |
| DZ male-female | 260 | .41 |

(Psychomotor retardation scores from Eaves et al., 1997)

Refined Matrix

$$\Omega = \begin{bmatrix} \sigma_g^2 + \sigma_c^2 + \sigma_e^2 & 2\phi\sigma_g^2 + \sigma_c^2 \\ 2\phi\sigma_g^2 + \sigma_c^2 & \sigma_g^2 + \sigma_c^2 + \sigma_e^2 \end{bmatrix}$$

Where,

ϕ is the kinship coefficient for the two individuals

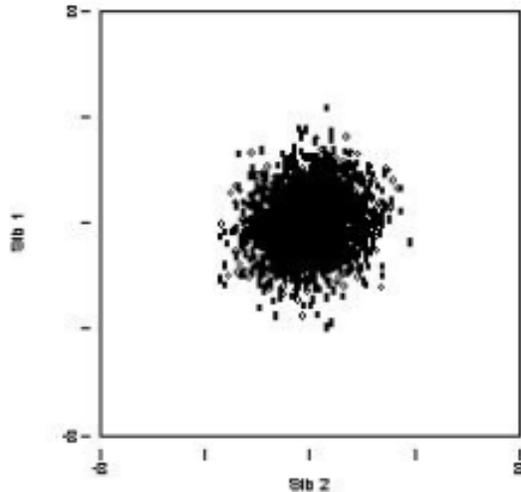
Interpretation...

- Fitting a maximum likelihood model...
 - Eaves et. al estimated (for males)
 - $\sigma_g^2 = .29$
 - $\sigma_c^2 = .24$
 - $\sigma_e^2 = .46$
 - Additive genetic effects could not explain similarities. Any idea why?

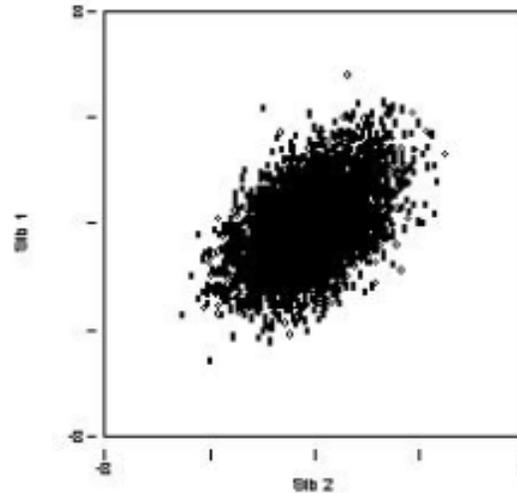
Incorporating IBD Coefficients

- Covariance might differ according to sharing at a particular locus
 - If locus contains genes that influence the trait
- Again, impose a genetic model and estimate model parameters

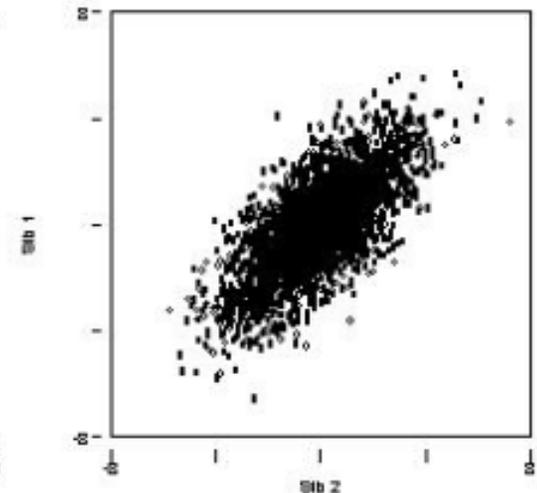
Linkage



IBD 0

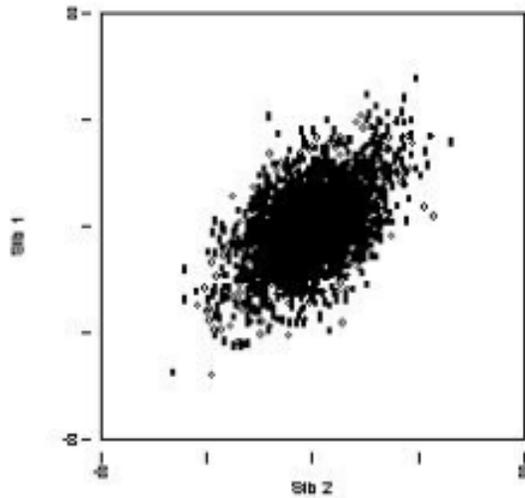


IBD 1

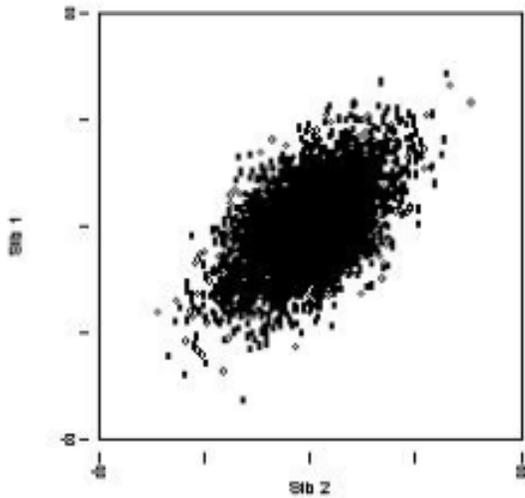


IBD 2

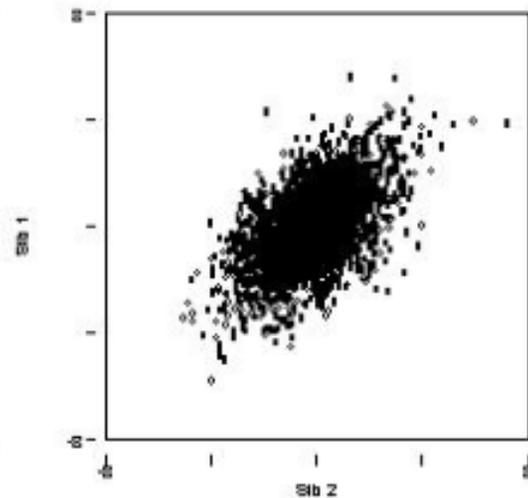
No Linkage



IBD 0



IBD 1



IBD 2

Relationship to IBD probabilities

- For non-inbred pair of relatives, marker or locus-specific kinship coefficients can be derived from IBD probabilities:

$$\varphi_{marker} = \frac{1}{4}P(IBM_{marker} = 1) + \frac{1}{2}P(IBM_{marker} = 2)$$

Variance-Covariance Matrix

$$\Omega = \begin{bmatrix} \sigma_a^2 + \sigma_g^2 + \sigma_e^2 & 2\varphi_{marker} \sigma_a^2 + 2\varphi \sigma_g^2 \\ 2\varphi_{marker} \sigma_a^2 + 2\varphi \sigma_g^2 & \sigma_a^2 + \sigma_g^2 + \sigma_e^2 \end{bmatrix}$$

Where,

φ is the kinship coefficient for the two individuals

φ_{marker} depends on the number of alleles shared IBD

Likelihood function, Incorporating Uncertain IBD

$$L = \prod_i \sum_{j=0,1,2} Z_{ij} (2\pi)^{-1} |\Omega_{IBD=j}|^{-1/2} e^{-1/2(\mathbf{y}-\boldsymbol{\mu})' \Omega_{IBD=j}^{-1} (\mathbf{y}-\boldsymbol{\mu})}$$
$$\approx \prod_i (2\pi)^{-1} |\Omega^*|^{-1/2} e^{-1/2(\mathbf{y}-\boldsymbol{\mu})' \Omega^{*-1} (\mathbf{y}-\boldsymbol{\mu})}$$

$Z_{ij} = P(IBD_i = j \mid \text{marker data})$ IBD sharing probabilities

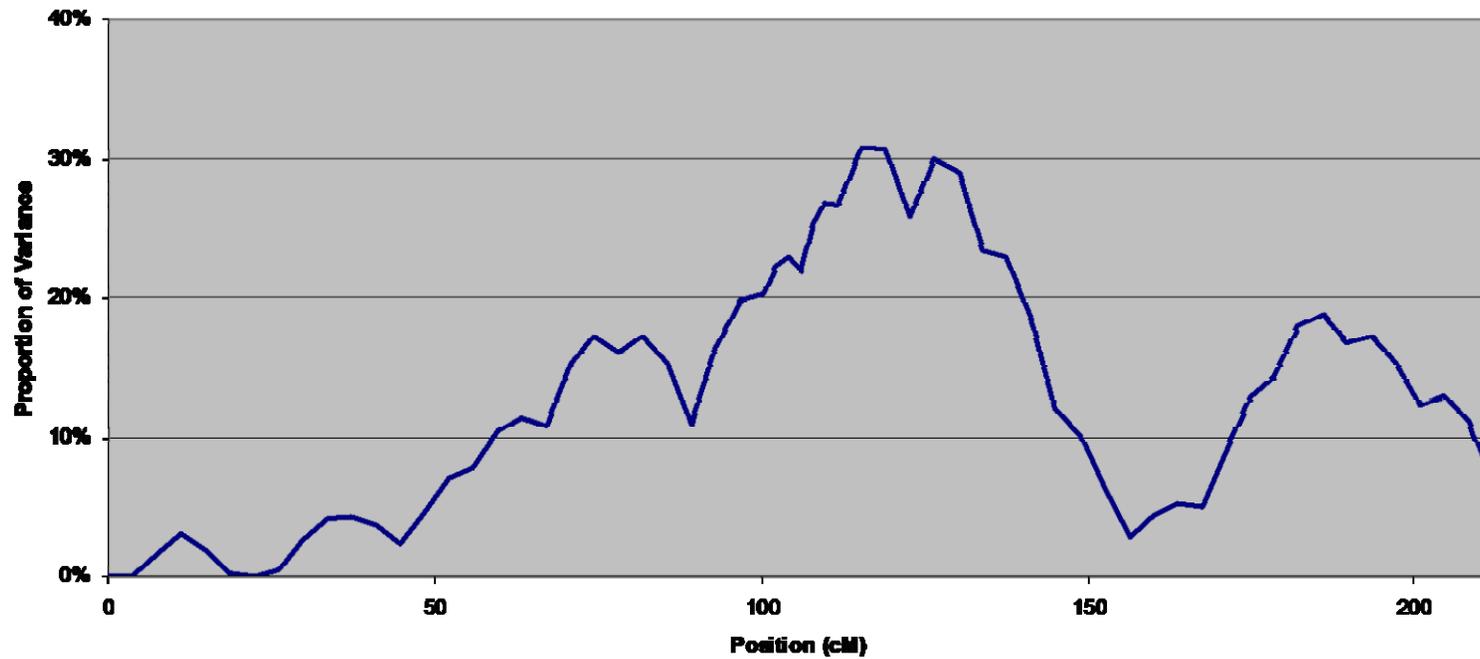
$\Omega^* = \sum_{j=0,1,2} Z_{ij} \Omega_{IBD=j}$ "Expected" Ω

How it works ...

- To find linkage to a particular trait...
- Collect sibling pair sample
- Calculate IBD for multiple points along genome
- Model covariance as a function of IBD sharing at each point

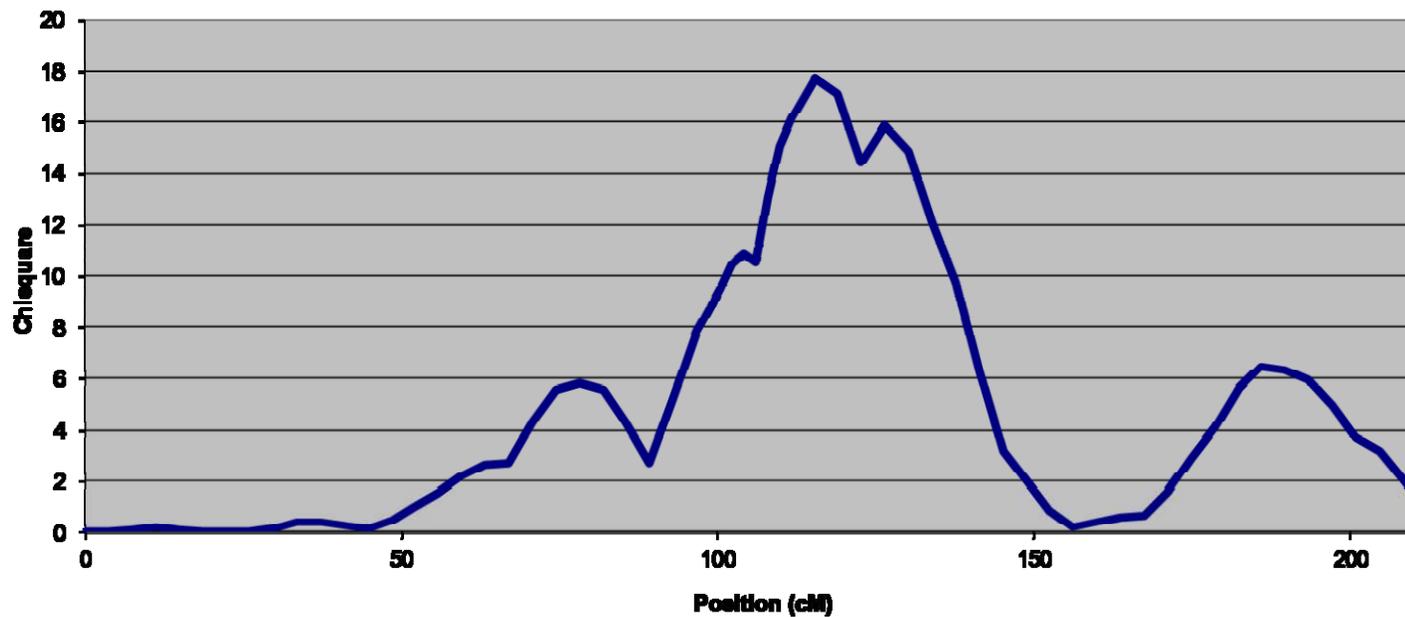
Example...

Estimated Major Gene Component



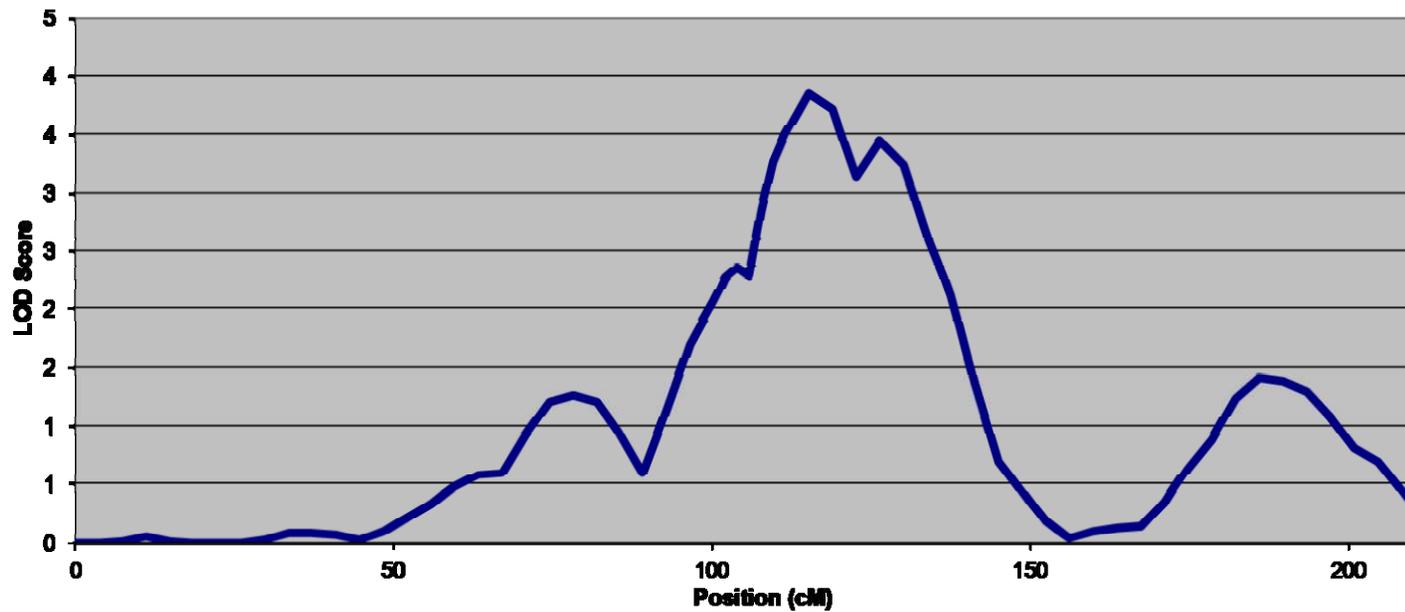
Example...

Likelihood Ratio Chisquare



Example...

LOD Score



So far ...

- Models for similarity between relative pairs
- Kinship coefficient used to estimate overall genetic effect
- Locus-specific coefficients used to detect genetic linkage

Extensions ...

- The model extends gracefully to other settings:
- For larger pedigrees, we extend the covariance matrix
- To model genetic association, we model specific means for each individual

Larger Pedigrees...

$$\Omega_{jk} = \begin{cases} \sigma_a^2 + \sigma_g^2 + \sigma_e^2 & \text{if } j = k \\ 2\varphi_{marker}\sigma_a^2 + 2\varphi\sigma_g^2 & \text{if } j \neq k \end{cases}$$

Where,

φ is the kinship coefficient for the two individuals

φ_{marker} depends on the number of alleles shared IBD

j and k index different individuals in the family

Multivariate density function

- Normal density function

$$L(y) = (2\pi)^{-1/2} \sigma^{-1} e^{-1/2(y-\mu)^2 / \sigma^2}$$

- Multivariate normal density function

$$L(\mathbf{y}) = 2\pi^{-n/2} |\mathbf{\Omega}|^{-1/2} e^{-1/2(\mathbf{y}-\boldsymbol{\mu})'\mathbf{\Omega}^{-1}(\mathbf{y}-\boldsymbol{\mu})}$$

- Extends univariate density function

Means Model

Expected Phenotype
for Individual i
(e.g. expected weight)

Estimated effects for covariates
(e.g. expected weight increases
1kg/year with age)

Measured Covariates
for Individual i
(e.g. age, sex, ...)

$$E(y_i) = \mu + \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_k x_{ik}$$

In addition to modeling variances and covariances, we can model the means

Simple Association Model

- Each copy of allele changes trait by a fixed amount
 - Include covariate counting copies for allele of interest
- Evidence for association when $a \neq 0$

$$E(y_i) = \mu + a * [\text{number of copies of mutant allele}]$$

$$E(y_i) = \mu + \beta_x X_i$$

X is the number of copies for allele of interest.

β_x is the estimated effect of each copy (the additive genetic value).

Today

- Analysis of quantitative traits
- Kinship coefficients
 - Measure of genetic similarity between two individuals
- Modeling covariance for pairs of individuals
 - estimating heritability
 - estimating locus-specific heritability
- Extending the model to larger pedigrees

Useful References

- Amos (1994)
Am J Hum Genet **54**:535-543
- Hopper and Matthews (1982)
Ann Hum Genet **46**:373–383
- Lange and Boehnke (1983)
Am J Med Genet **14**:513-24