# Variance Component Models 

## for Quantitative Traits

## Biostatistics 666

## Today

- Analysis of quantitative traits
- Kinship coefficients
- measure of genetic similarity between two individuals
- Modeling covariance for pairs of individuals
- estimating heritability
- estimating locus-specific heritability
- Extending the model to larger pedigrees


## Kinship Coefficients

Summarize genetic similarity between pairs of individuals.

In a variance components model, they predict the phenotypic similarity between individuals.

## Kinship Coefficients - Definition

- Given two individuals
- One with genes $\left(\mathrm{g}_{\mathrm{i}}, \mathrm{g}_{\mathrm{j}}\right)$
- The other with genes $\left(g_{k}, g_{l}\right)$
- The kinship coefficient is:
- $1 / 4 \mathrm{P}\left(\mathrm{g}_{\mathrm{i}} \equiv \mathrm{g}_{\mathrm{k}}\right)+1 / 4 \mathrm{P}\left(\mathrm{g}_{\mathrm{j}}=\mathrm{g}_{\mathrm{l}}\right)+1 / 4 \mathrm{P}\left(\mathrm{g}_{\mathrm{j}}=\mathrm{g}_{\mathrm{k}}\right)+1 / 4 \mathrm{P}\left(\mathrm{g}_{\mathrm{j}}=\mathrm{g}_{\mathrm{l}}\right)$
- where " $\equiv$ " represents identity by descent (IBD)
- Probability that alleles sampled at random from each individual are IBD


## Some kinship coefficients

Siblings ( $\varphi=1 / 4$ ) Half-Sibs ( $\varphi=1 / 8$ )

MZ Twins ( $\varphi=1 / 2$ )


Unrelated ( $\varphi=0$ )


## What about other relatives?

- For any two related individuals i and $\mathrm{j} . .$.
- ... use a recursive algorithm allows calculation of kinship coefficient
- Algorithm requires an order for individuals in the pedigree where ancestors precede descendants
- That is where for any $i>j, i$ is not ancestor of $j$
- Such an order always exists (e.g. the birth order!)


## Computing Kinship Coefficients

The recursive definition is then (for $\mathrm{i} \geq \mathrm{j}$ ):

$$
\varphi_{i j}=\left\{\begin{array}{cc}
0 & i \text { and } j \text { are founders } \\
1 / 2 & i=j, i \text { is a founder } \\
1 / 2\left(\varphi_{\text {mother }(i) j}+\varphi_{\text {father }(i) j}\right) & i \neq j \\
1 / 2\left(1+\varphi_{\text {mother }(i) \text { father }(i)}\right) & i=j
\end{array}\right.
$$

## An example pedigree...

- Can you find ...

Suitable ordering for recursive calculation?

Calculate kinship coefficient between shaded individuals?

## Inbreeding Coefficients

The kinship coefficient is related to the inbreeding coefficient

If $\phi_{i i}>0.5$, individual $i$ is inbred

The inbreeding coeffient is $\mathrm{f}_{\mathrm{i}}=\phi_{\text {mother }(i) \text { father }(i)}=2\left(\phi_{i i}-0.5\right)$

## So far

Summarize genetic similarity between any two individuals ...

Next, we will proceed to build a simple model for their phenotypes

## Simplest Data Structure

Pairs of related individuals

- Siblings (or twins!)
- Parent-Offspring

Corresponding phenotype measurements
${ }^{\circ} \mathrm{y}=\left(\mathrm{y}_{1}, \mathrm{y}_{2}\right)^{\prime}$

## Elements for a simple model

If the trait is normally distributed ...

- Model mean and variance for $\mathbf{y}_{1}$ and $\mathbf{y}_{2}$
- Mean and variance could be assumed equal ...
- ... or they could depend on some covariates
- But we are also interested in covariance between the two ...


## Variance-Covariance Matrix

$$
\Omega=\left[\begin{array}{cc}
V\left(y_{1}\right) & \operatorname{Cov}\left(y_{1}, y_{2}\right) \\
\operatorname{Cov}\left(y_{1}, y_{2}\right) & V\left(y_{2}\right)
\end{array}\right]
$$

Model must describe not only variance of each observation but also covariance for pairs of observations

## Bivariate density function

- Normal density function

$$
L(y)=\frac{1}{\sqrt{2 \pi}} \sigma^{-1} e^{-1 / 2(y-\mu)^{2} / \sigma^{2}}
$$

Bivariate normal density function

$$
L(\mathbf{y})=\frac{1}{2 \pi}|\Omega|^{-1 / 2} e^{-1 / 2(y-\mu) / \Omega^{-1}(y-\mu)}
$$

Extends univariate density function

## Intuition on Normal Densities



## Bivariate Normal Densities

## $L(\mathbf{y})=\left.(2 \pi)^{-1} \Omega\right|^{-1 / 2} e^{-1 / 2(\mathbf{y}-\mu)^{\prime} \Omega^{-1}(\mathbf{y}-\mu)}$ <br> Scaling parameter, penalizes settings with large variances <br> Distance between observation and its expected value

## Variability in Height, Independent Observations



## Variability in Height, Pairs of Observations

DZ twin Height $\mathrm{r}=0.32$

In a sample of twin or sibling pairs, we could use all the data to estimate means, variances and even covariances...

(Data from David Duffy)

## Height in DZ and MZ twins



(How would you interpret these data from David Duffy?)

## Incorporating Kinship Coeffic ients

- If genes influence trait ...
- Covariance will differ for each class of relative pair
- Instead of estimating covariance for each relationship, ...
- Impose genetic model that incorporates kinship and relates covariance between different classes of relative pair


## A Simple Model for the Variance-Covariance Matrix

$\Omega=\left[\begin{array}{cc}\sigma_{g}^{2}+\sigma_{e}^{2} & 2 \varphi \sigma_{g}^{2} \\ 2 \varphi \sigma_{g}^{2} & \sigma_{g}^{2}+\sigma_{e}^{2}\end{array}\right]$

Where,
$\varphi$ is the kinship coefficient for the two individuals

## Example...

|  | $\mathbf{N}$ | $\mathbf{r}$ |
| ---: | ---: | ---: |
| MZ males | 292 | .80 |
| MZ females | 380 | .80 |
| DZ males | 179 | .47 |
| DZ females | 184 | .55 |
| DZ male-female | 284 | .41 |

(Reading ability scores from Eaves et al., 1997)

## Interpretation...

Fitting a maximum likelihood model...
${ }^{-}$Eaves et. al estimated
${ }^{-} \sigma_{g}{ }^{2}=.81$

- $\sigma_{\mathrm{e}}{ }^{2}=.19$
- Found no evidence for sex differences
- Saturated model did not improve fit


## So far

Model allows us to estimate the genetic contribution to the variation in any trait

- Incorporates different relative pairs ...
- But it doesn't always fit...
- Fortunately, the model can be easily refined


## Another Example...

|  | $\mathbf{N}$ | $\mathbf{r}$ |
| ---: | ---: | ---: |
| MZ males | 271 | .56 |
| MZ females | 353 | .52 |
| DZ males | 167 | .33 |
| DZ females | 165 | .45 |
| DZ male-female | 260 | .41 |

(Psychomotor retardation scores from Eaves et al., 1997)

## Refined Matrix

$$
\Omega=\left[\begin{array}{cc}
\sigma_{g}^{2}+\sigma_{c}^{2}+\sigma_{e}^{2} & 2 \varphi \sigma_{g}^{2}+\sigma_{c}^{2} \\
2 \varphi \sigma_{g}^{2}+\sigma_{c}^{2} & \sigma_{g}^{2}+\sigma_{c}^{2}+\sigma_{e}^{2}
\end{array}\right]
$$

Where,
$\varphi$ is the kinship coefficient for the two individuals

## Interpretation...

Fitting a maximum likelihood model...

- Eaves et. al estimated (for males)
$-\sigma_{g}{ }^{2}=.29$
$-\sigma_{\mathrm{c}}{ }^{2}=.24$
- $\sigma_{\mathrm{e}}{ }^{2}=.46$
- Additive genetic effects could not explain similarities. Any idea why?


## Inc orporating IBD Coefficients

- IBD coefficients measure genetic similarity at a specific locus
- Related individuals might share 0, 1 or 2 alleles
- Covariance might differ according to sharing at a particular locus
- If locus contains genes that influence the trait
- Again, impose a genetic model and estimate model parameters


## Linkage





IBD 2

## No Linkage





## Relationship to IBD probabilities

For non-inbred pair of relatives, marker or locus-specific kinship coefficients can be derived from IBD probabilities:
$\varphi_{\text {marker }}=1 / 4 P\left(I B D_{\text {marker }}=1\right)+1 / 2 P\left(I B D_{\text {marker }}=2\right)$

## Variance-Covariance Matrix

$\Omega=\left[\begin{array}{cc}\sigma_{a}^{2}+\sigma_{g}^{2}+\sigma_{e}^{2} & 2 \varphi_{\text {marker }} \sigma_{a}^{2}+2 \varphi \sigma_{g}^{2} \\ 2 \varphi_{\text {marker }} \sigma_{a}^{2}+2 \varphi \sigma_{g}^{2} & \sigma_{a}^{2}+\sigma_{g}^{2}+\sigma_{e}^{2}\end{array}\right]$

Where,
$\varphi$ is the kinship coefficient for the two individuals
$\varphi_{\text {marker }}$ depends on the number of alleles shared IBD

## Likelihood function, Incorporating Uncertain IBD

$$
\begin{aligned}
L & =\prod_{i} \sum_{j=0,1,2 j} Z_{i j}(2 \pi)^{-1}\left|\Omega_{I B D=j}\right|^{-1 / 2} e^{-1 / 2(y-\mu) \Omega_{i D D}^{10}=(y-\mu)} \\
& \approx \prod_{i}(2 \pi)^{-1}\left|\Omega^{*}\right|^{-1 / 2} e^{-1 / 2(y-\mu) \Omega^{*+1}(y-\mu)}
\end{aligned}
$$

$Z_{i j}=P\left(I B D_{i}=j \mid\right.$ marker data $) \quad$ IBD sharing probabilities
$\Omega^{*}=\sum_{j=0,1,2} Z_{i j} \Omega_{I B D=j} \quad$ "Expected" $\Omega$

## How it works ...

- To find linkage to a particular trait...
- Collect sibling pair sample
- Calculate IBD for multiple points along genome
- Model covariance as a function of IBD sharing at each point


## Example...

Estimated Major Gene Component $\sigma_{\mathrm{a}}{ }^{2}$


## Example...

## Likelihood Ratio Chisquared

 $2 \ln L\left(\sigma_{a}^{2}\right) / L\left(\sigma_{a}^{2}=0\right)$

## Example...

## LOD Score <br> $\log _{10} L\left(\sigma_{a}^{2}\right) / L\left(\sigma_{a}^{2}=0\right)$



## So far

Models for similarity between relative pairs

Kinship coefficient used to estimate overall genetic effect

- Locus-specific coefficients used to detect genetic linkage


## Examples of Useful Extensions...

- For larger pedigrees, expand the covariance matrix
- For populations with individuals of uncertain relatedness, estimate kinship using genotypes
- To model genetic association, allow individual specific means to depend on genotype


## Larger Pedigrees...

$$
\Omega_{j k}=\left\{\begin{array}{cc}
\sigma_{a}^{2}+\sigma_{g}^{2}+\sigma_{e}^{2} & \text { if } j=k \\
2 \varphi_{\text {marker }} \sigma_{a}^{2}+2 \varphi \sigma_{g}^{2} & \text { if } j \neq k
\end{array}\right.
$$

Where,
$\varphi$ is the kinship coefficient for the two individuals $\varphi_{\text {marker }}$ depends on the number of alleles shared IBD
$j$ and $k$ index different individuals in the family

## Multivariate density function

Normal density function

$$
L(y)=(2 \pi)^{-1 / 2} \sigma^{-1} e^{-y / 2(y-\mu)^{2} / \sigma^{2}}
$$

- Multivariate normal density function

$$
L(\mathbf{y})=2 \pi^{-1 / 2}|\Omega|^{-1 / 2} e^{-1 / 2(y-\mu) \Omega^{-1}(y-\mu)}
$$

Extends univariate density function

## Covariate and Genotype effects

Expected Phenotype for Individual $i$
(e.g. expected weight)

Estimated effects for covariates (e.g. expected weight increases $1 \mathrm{~kg} / \mathrm{year}$ with age)

Measured Covariates for Individual $i$ (e.g. age, sex, genotype)

In addition to modeling variances and covariances, can model fixed effects

## Simple Association Model

- Each copy of allele changes trait by a fixed amount - Include covariate counting copies for allele of interest
- Evidence for association when $\mathrm{a} \neq 0$
$g_{i}=$ number of copies of allele of interest in individual $i$

$$
E\left(y_{i}\right)=\mu+\beta_{g} g_{i}
$$

$\beta_{g}$ is effect of each allele (the additive genetic value).

## Relatedness in Populations

- Although we have focused on individuals of known relationship, ...

Marker data can also be used to estimate relatedness.

For example, Kang et al (2010) use:

$$
\hat{\phi}_{i j}=\frac{1}{M} \sum_{m=1}^{M} \frac{\left(g_{i m}-2 p_{m}\right)\left(g_{j m}-2 p_{m}\right)}{4 p_{m}\left(1-p_{m}\right)}
$$

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## Useful References

- Amos (1994) Am J Hum Genet 54:535-543
- Hopper and Matthews (1982)

Ann Hum Genet 46:373-383

- Lange and Boehnke (1983)

Am J Med Genet 14:513-24

